

# ECON6190 Section 14

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2. [Hansen 14.4] You have the point estimate  $\hat{\theta} = 0.45$  and standard errors  $s(\hat{\theta}) = 0.28$ . You are interested in  $\beta = \exp(\theta)$ .

- Find  $\hat{\beta}$ .
- Use the delta method to find a standard error  $s(\hat{\beta})$ .
- Use the above to calculate a 95% asymptotic confidence interval for  $\hat{\beta}$ .
- Calculate a 95% asymptotic confidence interval  $[L, U]$  for the original parameter  $\theta$ . Calculate a 95% asymptotic confidence interval for  $\beta$  as  $[\exp(L), \exp(U)]$ . Can you explain why this is a valid choice? Compare this interval with your answer in (c).

(a) use plug-in estimator for  $\beta$

$$\hat{\beta} = \exp(\hat{\theta}) = \exp(0.45) \approx 1.57$$

(b) Assume  $\hat{\theta}$  is asymptotically normal, ie.  $\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{d} \mathcal{N}(0, \underline{V_{\theta}})$

asymptotic variance of  $\theta$

By delta method,

$$\begin{aligned} \sqrt{n}(\hat{\beta} - \beta) &= \sqrt{n}(\exp(\hat{\theta}) - \exp(\theta)) \\ &\stackrel{\text{by Taylor expansion}}{=} \sqrt{n} \exp(\tilde{\theta})(\hat{\theta} - \theta), \text{ where } \tilde{\theta} \text{ in between } \hat{\theta} \text{ and } \theta \\ &\xrightarrow{d} \mathcal{N}(0, \underbrace{(\exp(\theta))^2 V_{\theta}}_{= \exp(2\theta) V_{\theta}}) \end{aligned}$$

The standard error of  $\hat{\beta}$  is

$$\begin{aligned} s(\hat{\beta}) &= \sqrt{\underbrace{\exp(2\hat{\theta})}_{\substack{\text{comes from } \sqrt{n} \text{ rescaling} \\ s(\hat{\theta})}} \hat{V}_{\theta}} = \exp(\hat{\theta}) \sqrt{\frac{\hat{V}_{\theta}}{n}} = \exp(\hat{\theta}) s(\hat{\theta}) \\ &= \exp(0.45)(0.28) \approx 0.44 \end{aligned}$$

(c) from (b)  $\Rightarrow \sqrt{n}(\hat{\beta} - \beta) \xrightarrow{d} \mathcal{N}(0, \exp(2\theta)V_{\theta})$

$\Rightarrow$  construct asymptotic CI by asymptotic pivotal quantities

$$\frac{\hat{\beta} - \beta}{\sqrt{\frac{\exp(2\theta)V_{\theta}}{n}}} \xrightarrow{d} \mathcal{N}(0, 1) \quad \leftarrow \text{asymptotic pivot}$$

Denote  $Z_{1-0.025}$  as the  $(1-0.025)$ th quantile of  $\mathcal{N}(0, 1)$

Then as  $n \rightarrow \infty$ ,

$$P\left(-Z_{1-0.025} \leq \frac{\hat{\beta} - \beta}{\sqrt{\frac{\exp(2\theta)V_{\theta}}{n}}} \leq Z_{1-0.025}\right) \rightarrow 1-0.05$$

$$P\left(\hat{\beta} - Z_{1-0.025}(0.44) \leq \beta \leq \hat{\beta} + Z_{1-0.025}(0.44)\right) \rightarrow 0.95$$

$\Rightarrow$  A 95% asymptotic CI for  $\beta$  is

$$\left[\hat{\beta} - Z_{1-0.025}(0.44), \hat{\beta} + Z_{1-0.025}(0.44)\right].$$

or approximately  $[0.71, 2.43]$ .

(d) ① Construct asymptotic 95% CI for  $\theta$ .

$$\text{Similarly, } \left[\hat{\theta} - Z_{1-0.025}S(\hat{\theta}), \hat{\theta} + Z_{1-0.025}S(\hat{\theta})\right] \approx \left[-\overset{L}{0.099}, \overset{u}{0.999}\right]$$

②  $\Rightarrow$  An alternative CI for  $\beta$  is

$$[\exp(L), \exp(u)] \approx [0.91, 2.72]$$

This is a valid 95% asymptotic CI because

$$\begin{aligned} &P(\exp(L) \leq \beta \leq \exp(u)) \\ &= P(\exp(L) \leq \exp(\theta) \leq \exp(u)) \quad \text{exp(.) monotonically } \uparrow \\ &= P(L \leq \theta \leq u) \rightarrow 1 - \alpha = 95\%. \end{aligned}$$

$[\exp(L), \exp(u)] = [0.91, 2.72]$  is longer than  $[0.71, 2.43]$ .

4. Let the random variable  $X$  be normally distributed with mean  $\mu$  and variance 1. You are given a random sample of 16 observations.

- Construct a one sided 95% confidence interval for  $\mu$  that has form  $[\hat{L}, \infty)$  for some statistic  $\hat{L}$ .
- Construct a two sided 95% confidence interval for  $\mu$ .
- Show that the rejection of the null  $H_0 : \mu = 0$  against  $H_1 : \mu \neq 0$  with size 5% based on t test corresponds to the rejection of  $H_0 : \mu = 0$  when zero does not lie in the 95% confidence interval for  $\mu$  constructed in part (b).
- How would your answers be affected when you would not have known the variance of the random variable?

$$X \sim \mathcal{N}(\mu, 1) \quad n=16$$

(a) Denote the sample average  $\bar{X}_n$  as an estimator for  $\mu$ .

$$\frac{\bar{X}_n - \mu}{\sqrt{\frac{\sigma^2}{n}}} \sim \mathcal{N}(0, 1) \quad \Rightarrow \text{Construct CI using pivotal quantity.}$$

Note:  $Z_{1-0.025} = 1.65$

$$P\left(\frac{\bar{X}_n - \mu}{\sqrt{\frac{\sigma^2}{n}}} \leq 1.65\right) = 0.95$$

$$P\left(\frac{\bar{X}_n - \mu}{\sqrt{\frac{1}{16}}} \leq 1.65\right) = 0.95$$

$$\Rightarrow P\left(\mu \geq \bar{X}_n - \frac{1.65}{4}\right) = 0.95$$

A one sided 95% CI for  $\mu$  is  $\left[\bar{X}_n - \frac{1.65}{4}, \infty\right)$

(b) Similarly,  $\frac{\bar{X}_n - \mu}{1/4} \sim N(0,1)$ , and  $Z_{1-0.025} = 1.96$ .

$$P\left(-Z_{1-0.025} \leq \frac{\bar{X}_n - \mu}{1/4} \leq Z_{1-0.025}\right) = 0.95$$

$$P\left(\bar{X}_n - \frac{1.96}{4} \leq \mu \leq \bar{X}_n + \frac{1.96}{4}\right) = 0.95$$

$\Rightarrow$  A two sided 95% CI for  $\mu$  is  $\left[\bar{X}_n - \frac{1.96}{4}, \bar{X}_n + \frac{1.96}{4}\right]$ .

(c) WTS: the rejection rule is the same.

For the second part, 0 doesn't lie in 95% CI for  $\mu$  in (b)

$$\Rightarrow 0 \notin \left[\bar{X}_n - \frac{1.96}{4}, \bar{X}_n + \frac{1.96}{4}\right]$$

$$\Leftrightarrow 0 < \bar{X}_n - \frac{1.96}{4} \quad \text{or} \quad 0 > \bar{X}_n + \frac{1.96}{4}$$

$$\Leftrightarrow \frac{\bar{X}_n}{1/4} > 1.96 \quad \text{or} \quad \frac{\bar{X}_n}{1/4} < -1.96 \quad \dots\dots\dots ①$$

For the first part: conduct test  $H_0: \mu=0$ ,  $H_1: \mu \neq 0$ , with size 5%.

reject  $H_0$  under  $H_0$ .

Under  $H_0$ :  $T = \frac{\bar{X}_n - 0}{\sqrt{\frac{1}{16}}} \sim N(0,1)$ ,  $Z_{1-0.025} = 1.96$

reject if  $\frac{\bar{X}_n}{1/4} < -1.96$  or  $\frac{\bar{X}_n}{1/4} > 1.96 \quad \dots\dots\dots ②$

①② are the same.

(d) variance unknown.

If variance unknown, we can estimate the variance using

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2.$$

Hence, by results from sampling chapter, we get the following finite sample

distribution  $\frac{\bar{X}_n - \mu}{\sqrt{\frac{S^2}{n}}} \sim t_{n-1}$

$\Rightarrow$  By similar construction, 95% CI is

$$\left[ \bar{x}_n - t_{15, 1-0.025} \left( \frac{s}{4} \right), \bar{x}_n + t_{15, 1-0.025} \left( \frac{s}{4} \right) \right].$$

- (e) How would your answers be affected when you would not have known the variance of the random variable but the sample size is 100?

If  $n=100 \Rightarrow$  we have "large" sample,

$$\frac{\bar{x}_n - \mu}{\sqrt{s^2/n}} \xrightarrow{d} \mathcal{N}(0,1) \text{ when } n \text{ is large.}$$

asymptotic approximation.

A valid 95% CI is  $\left[ \bar{x}_n - \underbrace{1.96}_{Z_{1-0.025}} \left( \frac{s}{10} \right), \bar{x}_n + 1.96 \left( \frac{s}{10} \right) \right]$