

ECON 6190 Section 2

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4. [Hansen 6.6] Show that  $\mathbb{E}[s] \leq \sigma$ , where  $s = \sqrt{s^2}$  and  $s^2$  is the sample variance.

sample variance  $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_n)^2$  has the property that  $\mathbb{E}[s^2] = \sigma^2$ .

$\Rightarrow$  unbiased estimator of  $\sigma^2$

$$\sigma = \sqrt{\sigma^2} = \sqrt{\mathbb{E}[s^2]}$$

### General Jensen's inequality

- If  $g(\cdot) : \mathbb{R}^m \rightarrow \mathbb{R}$  is convex, then for any random vector  $X$  for which  $\mathbb{E}\|X\| < \infty$  and  $\mathbb{E}|g(X)| < \infty$

$$g(\mathbb{E}[X]) \leq \mathbb{E}[g(X)]$$

If  $g(\cdot) : \mathbb{R}^m \rightarrow \mathbb{R}$  is concave, then the inequality is reversed

- Jensen's inequality holds conditionally as well

By Jensen's inequality,  $g(x) = \sqrt{x}$  is concave.

$$g(\mathbb{E}[s^2]) \geq \mathbb{E}[g(s^2)]$$

$$\Leftrightarrow \underbrace{\sqrt{\mathbb{E}[s^2]}}_{\sigma} \geq \underbrace{\mathbb{E}[\sqrt{s^2}]}_{\mathbb{E}[s]}$$

$$\Rightarrow \mathbb{E}[s] \leq \sigma.$$

5. [Hong 3.31] Show that if  $X$  is a continuous random variable, then

$$\min_a \mathbb{E}|X - a| = \mathbb{E}|X - m|,$$

where  $m$  is the median of  $X$ .

$$\mathbb{E}[g(x)] = \int g(x) f(x) dx$$

$$\begin{aligned} \mathbb{E}[|x-a|] &= \int_{-\infty}^{\infty} |x-a| f(x) dx \\ &= \int_{-\infty}^a (a-x) f(x) dx + \int_a^{\infty} (x-a) f(x) dx \end{aligned}$$

$\underbrace{\hspace{10em}}_{\text{"}f(x,t)\text{"}}$

• Leibniz integral rule

Suppose  $g(x) = \int_{\alpha(x)}^{\beta(x)} f(x, t) dt$ , then

$$\frac{d}{dx} g(x) = \int_{\alpha(x)}^{\beta(x)} \frac{\partial f(x, t)}{\partial x} dt + \left( \frac{d}{dx} \beta(x) \right) f(x, \beta(x)) - \left( \frac{d}{dx} \alpha(x) \right) f(x, \alpha(x))$$

$x \rightarrow a$   
 $t \rightarrow x$

By Leibniz integral rule:

$$\begin{aligned} \frac{dE[|X-a|]}{da} &= \int_{-\infty}^a \underbrace{\frac{d(a-x)f(x)}{da}}_{f(x)} dx + \overbrace{\left( \frac{da}{da} \right) (a-a)f(a)}^{=0} - 0 \\ &\quad + \int_a^{\infty} \underbrace{\frac{d(x-a)f(x)}{da}}_{-f(x)} dx + 0 - \underbrace{\left( \frac{da}{da} \right) (a-a)f(a)}_{=0} \\ &= \int_{-\infty}^a f(x) dx - \int_a^{\infty} f(x) dx \quad (*) \end{aligned}$$

function of  $a$

Set  $(*) = 0 \Rightarrow \int_{-\infty}^a f(x) dx = \int_a^{\infty} f(x) dx$

$\Rightarrow P(X \leq a) = P(X > a)$

since  $P(X \leq a) + P(X > a) = 1 \Rightarrow P(X \leq a) = P(X > a) = 1/2$ .

By def of the median,  $a = m$ .

To check at  $a = m$  is a minimum,  $\epsilon > 0$ .

consider to the left of  $m$ ,  $\int_{-\infty}^{m-\epsilon} f(x) dx < 1/2$  and  $\int_{m-\epsilon}^{\infty} f(x) dx > 1/2 < 0$

consider to the right of  $m$ ,  $\int_{-\infty}^{m+\epsilon} f(x) dx > 1/2$  and  $\int_{m+\epsilon}^{\infty} f(x) dx < 1/2 > 0$

$\Rightarrow \min$



6. [Mid-term, Fall 2021] Let  $X$  be a random variable with conditional density

$$f(x|\theta) = \begin{cases} \frac{1}{\theta} & \text{if } 0 < x < \theta \\ 0 & \text{otherwise} \end{cases}.$$

Usually we treat parameter  $\theta$  as a constant. Now suppose  $\theta > 0$  is treated as a random variable with density

$$g(\theta) = \begin{cases} \theta e^{-\theta} & \text{if } \theta > 0 \\ 0 & \text{if } \theta \leq 0 \end{cases},$$

where we use notation  $\theta$  as both the random variable and the specific values it can take. Answer this following questions. (This question does not require any prior knowledge on Bayesian statistics, but is a test of your understanding of the key notions introduced in class.)

- (a) Find  $f(x)$ , the marginal density of  $X$ .  
 (b) Find  $g(\theta|x)$ , the conditional density of  $\theta$  given  $X = x$ .  
 (c) Find  $\mathbb{E}[(\theta - a)^2|X = x]$  for some given constant  $a$ . (You are NOT required to work out the final integration.)

(a) <sup>1)</sup> find joint distribution <sup>2)</sup> find marginal of the other variable

$$h(x, \theta) = f(x|\theta)g(\theta) = \begin{cases} \frac{1}{\theta} \theta e^{-\theta} = e^{-\theta}, & 0 < x < \theta \\ 0 & , \text{ o/w} \end{cases}$$

$$\begin{aligned} f(x) &= \int_{-\infty}^{\infty} h(x, \theta) d\theta = \int_x^{\infty} e^{-\theta} d\theta = -e^{-\theta} \Big|_x^{\infty} \\ &= 0 - (-e^{-x}) \\ &= e^{-x}, \quad x > 0 \end{aligned}$$

$$\Rightarrow f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0, & \text{o/w} \end{cases}$$

(b) For  $x > 0$ ,

$$g(\theta|x) = \frac{h(\theta, x)}{f(x)} = \begin{cases} \frac{e^{-\theta}}{e^{-x}} = e^{x-\theta}, & 0 < x < \theta \\ \frac{0}{e^{-x}} = 0, & x \geq \theta \end{cases}$$

$g(\theta|x)$  is not defined for  $x \leq 0$ .

$$\begin{aligned}
 c) \quad E[(\theta - a)^2 | X=x] &= \int (\theta - a)^2 g(\theta|x) d\theta \\
 &= \int_x^\infty (\theta - a)^2 e^{x-\theta} d\theta \\
 &= e^x \int_x^\infty (\theta - a)^2 e^{-\theta} d\theta.
 \end{aligned}$$

function of  $\theta$   
 $\rightarrow g(\theta|x)$

4. [Hong 5.47] Suppose  $X$  and  $Y$  are random variables such that  $\mathbb{E}[Y|X] = 7 - (1/4)X$  and  $\mathbb{E}[X|Y] = 10 - Y$ . Determine the correlation between  $X$  and  $Y$ .

Recall:  $\text{Corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{var}(X)\text{var}(Y)}}$

$$\begin{aligned}
 \text{cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\
 &= E[XY] - E[X]E[Y]
 \end{aligned}$$

$$E[X] \stackrel{\text{LIE}}{=} E[E[X|Y]] = E[10 - Y] = 10 - E[Y] \quad (1)$$

$$E[Y] = E[E[Y|X]] = E[7 - 1/4 X] = 7 - 1/4 E[X] \quad (2)$$

Combine (1) (2):

$$\begin{cases} E[X] = 10 - E[Y] \\ E[Y] = 7 - 1/4 E[X] \end{cases} \Rightarrow \begin{cases} E[X] = 4 \\ E[Y] = 6 \end{cases}$$

$$\begin{aligned}
 E[XY] &\stackrel{\text{LIE}}{=} E[E[XY|X]] = E[X E[Y|X]] \\
 &= E[X(7 - 1/4 X)] \\
 &= 7E[X] - 1/4 E[X^2] \\
 &= 28 - 1/4 E[X^2]
 \end{aligned}$$

$$\begin{aligned}
 E[XY] &= E[E[X|Y]Y] = E[Y E[X|Y]] \\
 &= E[Y(10 - Y)] \\
 &= 10E[Y] - E[Y^2] \\
 &= 60 - E[Y^2]
 \end{aligned}$$

$$\Rightarrow \begin{cases} E[X^2] = 112 - 4E[XY] \\ E[Y^2] = 60 - E[XY] \end{cases}$$

$$\text{var}(X) = E[X^2] - \underbrace{(E[X])^2}_4 = 112 - 4E[XY] - 16 = 4(24 - E[XY])$$

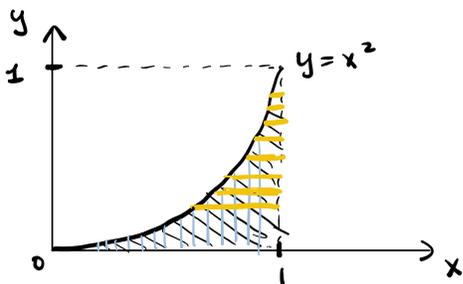
$$\text{var}(Y) = E[Y^2] - (E[Y])^2 = 60 - E[XY] - 36 = 24 - E[XY]$$

$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{var}(X)\text{var}(Y)}} = \frac{E[XY] - 24}{\sqrt{4(24 - E[XY])(24 - E[XY])}} = -1/2$$

2. [Hong 5.4]

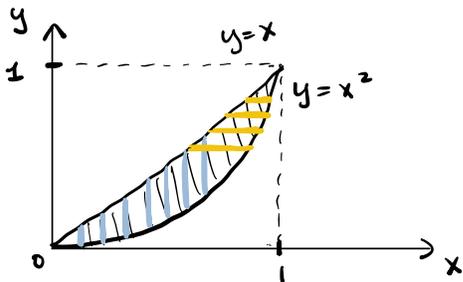
(a) Find  $P(X > \sqrt{Y})$  if  $X$  and  $Y$  are jointly distributed with pdf  $f_{X,Y}(x, y) = x + y$  for  $0 \leq x \leq 1, 0 \leq y \leq 1$ .

(b) Find  $P(X^2 < Y < X)$  if  $X$  and  $Y$  are jointly distributed with pdf  $f_{X,Y}(x, y) = 2x$  for  $0 \leq x \leq 1, 0 \leq y \leq 1$ .



$$\begin{aligned} P(X > \sqrt{Y}) &= \iint_A f_{X,Y}(x, y) dy dx \\ &= \int_0^1 \int_0^{x^2} f_{X,Y}(x, y) dy dx = \int_0^1 \int_{y^2}^1 f_{X,Y}(x, y) dx dy \\ &= 7/20 \end{aligned}$$

$x^2 > y$   
 $y < x^2$



$$\begin{aligned} P(X^2 < Y < X) &= \iint_A 2x dy dx \\ &= \int_0^1 \int_{x^2}^x 2x dy dx = \int_0^1 \int_y^{\sqrt{y}} 2x dx dy \\ &= 1/6 \end{aligned}$$