

ECON6190 Section 14

Yiwei Sun

Dec. 6, 2024

2. [Hansen 14.4] You have the point estimate $\hat{\theta} = 0.45$ and standard errors $s(\hat{\theta}) = 0.28$. You are interested in $\beta = \exp(\theta)$.

- (a) Find $\hat{\beta}$.
- (b) Use the delta method to find a standard error $s(\hat{\beta})$.
- (c) Use the above to calculate a 95% asymptotic confidence interval for $\hat{\beta}$.
- (d) Calculate a 95% asymptotic confidence interval $[L, U]$ for the original parameter θ . Calculate a 95% asymptotic confidence interval for β as $[\exp(L), \exp(U)]$. Can you explain why this is a valid choice? Compare this interval with your answer in (c).

(a) use plug-in estimator for β

$$\hat{\beta} = \exp(\hat{\theta}) = \exp(0.45) \approx 1.57$$

(b) Assume $\hat{\theta}$ is asymptotically normal, i.e. $\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{d} \mathcal{N}(0, V_{\theta})$

asymptotic variance of θ

By delta method,

$$\begin{aligned} \sqrt{n}(\hat{\beta} - \beta) &= \sqrt{n}(\exp(\hat{\theta}) - \exp(\theta)) \\ &\stackrel{\text{by Taylor expansion}}{=} \sqrt{n} \exp(\tilde{\theta})(\hat{\theta} - \theta), \text{ where } \tilde{\theta} \text{ in between } \hat{\theta} \text{ and } \theta \\ &\xrightarrow{d} \mathcal{N}(0, \underbrace{(\exp(\theta))^2}_{= \exp(2\theta)} V_{\theta}) \end{aligned}$$

The standard error of $\hat{\beta}$ is

$$\begin{aligned} s(\hat{\beta}) &= \sqrt{\frac{\exp(2\hat{\theta}) \hat{V}_{\theta}}{n}} = \exp(\hat{\theta}) \sqrt{\frac{\hat{V}_{\theta}}{n}} = \exp(\hat{\theta}) s(\hat{\theta}) \\ &\stackrel{\text{comes from } \sqrt{n} \text{ rescaling}}{=} s(\hat{\theta}) = \exp(0.45)(0.28) \approx 0.44 \end{aligned}$$

(c) from (b) $\Rightarrow \sqrt{n}(\hat{\beta} - \beta) \xrightarrow{d} \mathcal{N}(0, \exp(2\theta)V_{\theta})$

\Rightarrow construct asymptotic CI by asymptotic pivotal quantities

$$\frac{\hat{\beta} - \beta}{\sqrt{\frac{\exp(2\theta)V_{\theta}}{n}}} \xrightarrow{d} \mathcal{N}(0, 1) \quad \leftarrow \text{asymptotic pivot}$$

Denote $Z_{1-0.025}$ as the $(1-0.025)$ th quantile of $\mathcal{N}(0, 1)$

Then as $n \rightarrow \infty$,

$$P\left(-Z_{1-0.025} \leq \frac{\hat{\beta} - \beta}{\sqrt{\frac{\exp(2\theta)V_{\theta}}{n}}} \leq Z_{1-0.025}\right) \rightarrow 1-0.05$$

$$P\left(\hat{\beta} - Z_{1-0.025}(0.44) \leq \beta \leq \hat{\beta} + Z_{1-0.025}(0.44)\right) \rightarrow 0.95$$

\Rightarrow A 95% asymptotic CI for β is

$$\left[\hat{\beta} - Z_{1-0.025}(0.44), \hat{\beta} + Z_{1-0.025}(0.44)\right].$$

or approximately $[0.71, 2.43]$.

(d) ① Construct asymptotic 95% CI for θ .

$$\text{Similarly, } \left[\hat{\theta} - Z_{1-0.025}S(\hat{\theta}), \hat{\theta} + Z_{1-0.025}S(\hat{\theta})\right] \approx \left[-\overset{L}{0.099}, \overset{U}{0.999}\right]$$

② \Rightarrow An alternative CI for β is

$$[\exp(L), \exp(U)] \approx [0.91, 2.72]$$

This is a valid 95% asymptotic CI because

$$\begin{aligned} &P(\exp(L) \leq \beta \leq \exp(U)) \\ &= P(\exp(L) \leq \exp(\theta) \leq \exp(U)) \quad \text{exp(.) monotonically } \uparrow \\ &= P(L \leq \theta \leq U) \rightarrow 1 - \alpha = 95\%. \end{aligned}$$

$[\exp(L), \exp(U)] = [0.91, 2.72]$ is longer than $[0.71, 2.43]$.

4. Let the random variable X be normally distributed with mean μ and variance 1. You are given a random sample of 16 observations.

- Construct a one sided 95% confidence interval for μ that has form $[\hat{L}, \infty)$ for some statistic \hat{L} .
- Construct a two sided 95% confidence interval for μ .
- Show that the rejection of the null $H_0 : \mu = 0$ against $H_1 : \mu \neq 0$ with size 5% based on t test corresponds to the rejection of $H_0 : \mu = 0$ when zero does not lie in the 95% confidence interval for μ constructed in part (b).
- How would your answers be affected when you would not have known the variance of the random variable?

$$X \sim \mathcal{N}(\mu, 1) \quad n=16$$

(a) Denote the sample average \bar{X}_n as an estimator for μ .

$$\frac{\bar{X}_n - \mu}{\sqrt{\frac{\sigma^2}{n}}} \sim \mathcal{N}(0,1) \quad \Rightarrow \text{Construct CI using pivotal quantity.}$$

Note: $Z_{1-0.025} = 1.65$

$$P\left(\frac{\bar{X}_n - \mu}{\sqrt{\frac{\sigma^2}{n}}} \leq 1.65\right) = 0.95$$

$$P\left(\frac{\bar{X}_n - \mu}{\sqrt{\frac{1}{16}}} \leq 1.65\right) = 0.95$$

$$\Rightarrow P\left(\mu \geq \bar{X}_n - \frac{1.65}{4}\right) = 0.95$$

A one sided 95% CI for μ is $\left[\bar{X}_n - \frac{1.65}{4}, \infty\right)$

(b) Similarly, $\frac{\bar{X}_n - \mu}{1/4} \sim \mathcal{N}(0,1)$, and $Z_{1-0.025} = 1.96$.

$$P\left(-Z_{1-0.025} \leq \frac{\bar{X}_n - \mu}{1/4} \leq Z_{1-0.025}\right) = 0.95$$

$$P\left(\bar{X}_n - \frac{1.96}{4} \leq \mu \leq \bar{X}_n + \frac{1.96}{4}\right) = 0.95$$

\Rightarrow A two sided 95% CI for μ is $\left[\bar{X}_n - \frac{1.96}{4}, \bar{X}_n + \frac{1.96}{4}\right]$.

(c) WTS: the rejection rule is the same.

For the second part, 0 doesn't lie in 95% CI for μ in (b)

$$\Rightarrow 0 \notin \left[\bar{X}_n - \frac{1.96}{4}, \bar{X}_n + \frac{1.96}{4}\right]$$

$$\Leftrightarrow 0 < \bar{X}_n - \frac{1.96}{4} \quad \text{or} \quad 0 > \bar{X}_n + \frac{1.96}{4}$$

$$\Leftrightarrow \frac{\bar{X}_n}{1/4} > 1.96 \quad \text{or} \quad \frac{\bar{X}_n}{1/4} < -1.96 \quad \dots \textcircled{1}$$

For the first part: conduct test $H_0: \mu=0$, $H_1: \mu \neq 0$, with size 5%.

reject H_0 under H_0 .

Under H_0 : $T = \frac{\bar{X}_n - 0}{\sqrt{\frac{1}{16}}} \sim \mathcal{N}(0,1)$, $Z_{1-0.025} = 1.96$

reject if $\frac{\bar{X}_n}{1/4} < -1.96$ or $\frac{\bar{X}_n}{1/4} > 1.96 \quad \dots \textcircled{2}$

$\textcircled{1} \textcircled{2}$ are the same.

(d) variance unknown.

If variance unknown, we can estimate the variance using

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2.$$

Hence, by results from sampling chapter, we get the following finite sample

distribution $\frac{\bar{X}_n - \mu}{\sqrt{\frac{S^2}{n}}} \sim t_{n-1}$

\Rightarrow By similar construction, 95% CI is

$$\left[\bar{x}_n - t_{\lfloor 1-0.025 \rfloor} \left(\frac{s}{\sqrt{n}} \right), \bar{x}_n + t_{\lfloor 1-0.025 \rfloor} \left(\frac{s}{\sqrt{n}} \right) \right].$$

- (e) How would your answers be affected when you would not have known the variance of the random variable but the sample size is 100?

If $n = 100 \Rightarrow$ we have "large" sample,

$$\frac{\bar{x}_n - \mu}{\sqrt{s^2/n}} \xrightarrow{d} \mathcal{N}(0,1) \text{ when } n \text{ is large.}$$

asymptotic approximation.

A valid 95% CI is $\left[\bar{x}_n - \underbrace{1.96}_{Z_{1-0.025}} \left(\frac{s}{\sqrt{n}} \right), \bar{x}_n + 1.96 \left(\frac{s}{\sqrt{n}} \right) \right]$